# Network Reconstruction via High-Dimensional ODEs 

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Joint work with Shizhe Chen \& Daniela Witten

## Nonlinear Dynamics in Gene Regulatory Networks

$$
\begin{aligned}
& \frac{\mathbf{a}}{\frac{d\left(\text { gene }_{3}\right)}{d t}=k_{1, s} \cdot \frac{1}{1+k_{1,3} \cdot \text { gene }_{3}}-k_{1, d} \cdot \text { gene }_{1}} \\
& \frac{d\left(\text { gene }_{2}\right)}{d t}=k_{2, s} \cdot \frac{k_{21} \cdot \text { gene }_{1}}{1+k_{2,1} \cdot \text { gene }_{1}}-k_{2, \mathrm{~d}} \cdot \text { gene }_{2} \\
& \frac{d\left(\text { gene }_{3}\right)}{d t}=k_{3, s} \cdot \frac{k_{3,1} \cdot \text { gene }_{1} \cdot k_{3,2} \cdot \text { gene }_{2}}{\left(1+k_{3,1} \cdot \text { gene }\right) \cdot\left(1+k_{3,2} \cdot \text { gene }_{2}\right)}-k_{3, d} \cdot \text { gene }_{3}
\end{aligned}
$$

b



Karlebach \& Shamir (2008), Nature Reviews: Molecular Cell Biology

## Time-course gene expression data



|  | 0 | 0.02 | 0.04 | 0.06 | 0.08 | 0.1 | 0.12 | 0.14 | 0.16 | 0.1 |
| :--- | ---: | :--- | :--- | :--- | :--- | :--- | :--- | :--- | :--- | :--- |
| rseB | 0.22 |  |  |  |  |  |  |  |  |  |
| cpxR | 0.99 |  |  |  |  |  |  |  |  |  |
| motB | 0.71 |  |  |  |  |  |  |  |  |  |
| mdtA | 0.62 |  |  |  |  |  |  |  |  |  |
| ompC | 0.44 |  |  |  |  |  |  |  |  |  |
| infB | 1.04 |  |  |  |  |  |  |  |  |  |



## E. coli Gene Regulatory Network ${ }^{1}$



In this graph: $\quad X_{1} \longrightarrow X_{2} \Longleftrightarrow X_{1}$ regulates $X_{2}$

[^0]
## Gene Regulatory Network as a System of ODEs

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The change in expression of one gene is "regulated" by the expressions of others at the same time point

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X_{j}^{\prime}(t)=f_{j}\left(X(t), \theta_{j}\right)
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Example: Linear ODEs, $X^{\prime}(t)=\Theta X(t) \quad(\Rightarrow X(t)=\exp (\Theta t) X(0))$

$$
\Theta=\left(\begin{array}{ccc}
-0.59 & -1.36 & 1.32 \\
0.00 & 1.18 & 0.62 \\
-1.52 & -0.93 & 0.00
\end{array}\right)
$$




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Gold Standard:

- Find $\theta$ such that $X(\cdot ; \theta)$ solves the ODE:

$$
\begin{array}{r}
\hat{\theta}=\arg \min _{\theta} \sum_{i=1}^{n}\left\|Y_{i}-X\left(t_{i} ; \theta\right)\right\|^{2} \\
\text { s.t. } X^{\prime}(t ; \theta)=f(X(t ; \theta), \theta)
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- Accurate but slow
- requires numerical solution of ODE for every candidate $\theta$
- $\sqrt{n}$-consistent ${ }^{2}$


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Collocation Methods ${ }^{3}$ :

- Two-stage estimation strategy:
- estimate $\hat{X}(t)$ and $\hat{X}^{\prime}(t)$ from data (e.g. kernel smoothing)
- find $\theta$ that minimizes deviation from ODE:

$$
\hat{\theta}=\arg \min _{\theta} \sum_{i=1}^{n}\left\|\hat{X}^{\prime}\left(t_{i}\right)-f\left(\hat{X}\left(t_{i}\right) ; \theta\right)\right\|_{2}^{2}
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- Fast but not exact
- easy computation
- not solving the ODE exactly!
${ }^{3}$ Varah (1982)


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- Exact form of $f$ not known!

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Solution:

- Assume that $f_{j}$ is additive

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- $X_{k} \longrightarrow X_{j}$ iff $f_{j k} \neq 0$


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- Calculate $\hat{X}_{j}^{\prime}(t) \equiv \partial \hat{X}_{j} / \partial t$ (similar to collocation)
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- For a truncated basis $\psi=\left(\psi_{1}, \ldots, \psi_{M}\right)^{\mathrm{T}}$,

$$
f_{j k}(\cdot)=\boldsymbol{\psi}(\cdot)^{\mathrm{T}} \theta_{j k}+\delta_{j k}(\cdot) \quad(\text { allowing } M \rightarrow \infty \text { with } n)
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\int_{0}^{1}\{\hat{X}_{j}^{\prime}(t)-\underbrace{\theta_{j 0}-\sum_{k=1}^{p} \psi\left(\hat{X}_{k}(t)\right)^{\mathrm{T}} \theta_{j k}}_{\hat{f}_{j}}\}^{2} d t+\lambda \sum_{k=1}^{p} \underbrace{\left[\int_{0}^{1}\left\{\psi\left(\hat{X}_{k}(t)\right)^{\mathrm{T}} \theta_{j k}\right\}^{2} d t\right]^{1 / 2}}_{\text {group lasso penalty }}
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A penalized regression problem!
${ }^{4}$ Wu et al (2014), Henderson \& Michailidis (2014)

Key Observation - I

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- Estimating the derivative is inefficient ${ }^{5}$ !
- Hard to pick optimal bandwidth for estimating $d \hat{X} / d t$ !

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X_{j}^{\prime}(t) \approx \theta_{j 0}+\sum_{k=1}^{p} \psi\left(X_{k}(t)\right)^{\mathrm{T}} \theta_{j k}
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- Integrating both sides, we get

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X_{j}(t)-X(0) \approx t \theta_{j 0}+\sum_{k=1}^{p}\left(\int_{0}^{t} \psi\left(X_{k}(s)\right)^{\mathrm{T}} d s\right) \theta_{j k}
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- Let $\Psi_{i k}=\int_{0}^{t_{i}} \psi\left(X_{k}(s)\right) d s$,

$$
Y_{i j} \approx X(0)+t_{i} \theta_{j 0}+\sum_{k=1}^{p} \Psi_{i k}^{\mathrm{T}} \theta_{j k}+\varepsilon_{i j}
$$

- $\Psi_{i k}$ can be estimated as $\int_{0}^{t_{i}} \psi\left(\hat{X}_{k}(s)\right) d s$

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- Find the minimizer $\hat{\theta}_{j}$ of

$$
\sum_{i=1}^{n}[Y_{i j}-\underbrace{t_{i} \theta_{j 0}-\sum_{i=1}^{n} \hat{\Psi}_{i k}^{T} \theta_{j k}}_{\hat{f}_{j}}]^{2}+\lambda \sum_{k=1}^{p} \underbrace{\left.\sum_{i=1}^{n}\left(\hat{\Psi}_{i k}^{T} \theta_{j k}\right)^{2}\right]^{1 / 2}}_{\text {group lasso }}
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whose theoretical properties are well-understood

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- With $\hat{\Psi}_{i k}$ we have an errors-in-variables regression
- Need a bound on || $\hat{\Psi}-\Psi \mid$

Theory - II

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- We establish a new concentration inequality to bound

$$
\int_{0}^{1}\left\{\hat{X}_{j}(t)-X_{j}(t)\right\}^{2} d t
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- This inequality allows us to bound $\|\hat{\Psi}-\Psi\|$ in high dimensions, when $\log p / n^{\alpha}=o(1)$ for some $0<\alpha<0.5$
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- Using this inequality, the bound for derivative is asymptotically worst
- We show that GRADE can consistently select the parents of each node in a sparse high-dimensional ODE network
- The proof requires establishing model selection consistency of (standardized) group lasso regression with errors-in-variables ${ }^{6}$

[^3]
## Simulation: Design



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## Simulation: Results

- NeRDS: Network Reconstruction via Dynamic Systems ${ }^{7}$
- GRADE


[^4]
## Application: DREAM-3 Challenge ${ }^{8}$ (Regulatory Networks)

- in silico data from 5 regulatory networks with $p=10$ or $100 ; n=50$
- A difficult task: non-additive ODEs with unobserved latent variables


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Table: Area Under ROC Curves for NeRDS and GRADE

|  | $p=10$ |  | $p=100$ |  |
| :---: | :---: | :---: | :---: | :---: |
| Network | NeRDS | GRADE | NeRDS | GRADE |
| Ecoli1 | 0.450 | $\mathbf{0 . 5 4 5}$ | 0.624 | $\mathbf{0 . 6 7 0}$ |
| Ecoli2 | 0.512 | $\mathbf{0 . 6 4 3}$ | 0.637 | $\mathbf{0 . 6 5 3}$ |
| Yeast1 | 0.486 | $\mathbf{0 . 6 7 9}$ | 0.610 | $\mathbf{0 . 6 3 6}$ |
| Yeast2 | 0.525 | $\mathbf{0 . 6 0 7}$ | 0.568 | $\mathbf{0 . 5 8 4}$ |
| Yeast3 | 0.467 | $\mathbf{0 . 5 7 6}$ | $\mathbf{0 . 6 1 7}$ | 0.567 |

[^5]
## Application: Brain Functional Connectivity

- Cortical activity map (CAM) project - Allen Institute for Brain Science
- Calcium fluorescent imaging in a region of visual cortex at 175 mm depth measured using two-photons technology
- 575 neurons $\rightarrow 25$ neuronal populations ( $5 \times 5$ grids with $\sim 20$ neurons)
- 3 stimuli: frequencies of 1,2 , and 4 Hz , at a $90^{\circ}$ spatial orientation
- $R=15$ repetitions, $n=60$ time points per repetition


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- More similar connectivity networks for closer frequencies


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- It uses the linearity in parameters of truncated bases to avoid the estimation of derivatives
- Empirical \& theoretical evidence shows improved performance
- GCV for bandwidth selection results in consistent estimates


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- Empirical \& theoretical evidence shows improved performance
- GCV for bandwidth selection results in consistent estimates
- How can this idea be generalize to non-additive ODEs?


## Acknowledgments:

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## Comparison with Methods for Linear ODEs

- Linear ODEs of the form $X^{\prime}(t)=\Theta X(t)+C$
- Comparison of GRADE with Hall \& Ma (2014) and Brunel et al (2014)


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## A Concentration Inequality

## Theorem:

Under standard assumptions and if $\varepsilon_{j}, j=1, \ldots, p$ are i.i.d. $N(0,1)$, the local polynomial regression estimator $\hat{X}(\cdot)$ satisfies

$$
\int_{0}^{1}\left\{\hat{X}_{j}(t)-X_{j}(t)\right\}^{2} d t \leq c_{1} n^{\frac{2 \beta}{2 \beta+1}(\alpha-0.5)}
$$

for all $j=1, \ldots, p$, with probability converging to 1 if

$$
p \exp \left(-c_{2} n^{2 \alpha}\right)=o(1)
$$

Remarks:

- Here, $\beta$ and $\alpha$ are constants related to the smoothness of $X$ and the choice of bandwidth for $\hat{X}$
- GCV, CV, and other methods can be used to choose the bandwidth
- For $\left\|\hat{X}^{\prime}-X^{\prime}\right\|_{2}$, the rate is $n^{\frac{2 \beta-2}{2 \beta-1}(\alpha-0.5)}$


## Model Selection Consistency

Theorem:
Let

$$
N_{j}^{*}=\left\{k:\left\|\theta_{j k}\right\|_{2} \neq 0\right\}, \quad j=1, \ldots, p
$$

be the true parents of $X_{j}(\cdot)$ in the ODE network, and $\hat{N}_{j}$ be its estimator using the proposed method. Then, under certain regularity conditions, as the number of time points $n$ increases,

$$
\mathbb{P}\left(\hat{N}_{j}=N_{j}^{*} \text { for all } j=1, \ldots, p\right) \rightarrow 1 .
$$

Remark:

- The proof requires establishing model selection consistency of (standardized) group lasso regression with errors-in-variables


[^0]:    ${ }^{1}$ Subnetwork of E-coli regulatory network (GeneNetWeaver, Schaffter et al, 2011)

[^1]:    ${ }^{3}$ Varah (1982)

[^2]:    ${ }^{5}$ More later...

[^3]:    ${ }^{6}$ Extending lasso (Loh \& Wainwright, 2012 and Rosenbaum \& Tsybakov, 2010)

[^4]:    ${ }^{7}$ Henderson \& Michailidis (2014)

[^5]:    ${ }^{8}$ Schaffter et al (2011)

