# **Shapley Effects** for Sensitivity Analysis with Dependent Inputs

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In this talk, we consider

$$f: \left\{ \begin{array}{l} \mathcal{X} = \mathcal{X}_1 \times \dots \mathcal{X}_d & \to & \mathcal{Y} \\ \mathbf{x} = (x_1, \dots, x_d) & \mapsto & \mathbf{y} = f(\mathbf{x}) \end{array} \right.$$

with

- *f* : mathematical or numerical model,
- x : uncertain input parameters,
- y : output.

We model the uncertainty on the input parameters by a probability distribution P on  $\mathcal{X}$  and get

$$Y = f(X_1, \ldots, X_d)$$

with the vector  $\mathbf{X} = (X_1, \dots, X_d)$  distributed as P.



Independent framework:  $P(d\mathbf{x}) = P_1(d\mathbf{x}_1) \dots P_d(d\mathbf{x}_d)$ 

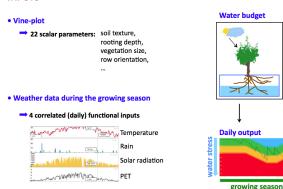
Why is the independent framework not always the right one?

Let us consider the following example: an agro-climatic model for the water status management of vineyard. Joint study with INRA and iTK (Montpellier, FRANCE).



**Project objective:** control of grape/wine quality. SA as decision support.

#### **INPUTS**



The soil texture was initially described by 3 scalar parameters: the percentages of argil, sand and silt.

These parameters are not independent as

% argil + % sand + % silt = 
$$100\%$$
.

In the study, this set of parameters has been replaced by a unique parameter aSoil describing the influence of the soil texture on its evaporation capacity.

Daily precipitations, solar radiation, mean air temperature and potential evapotranspiration are temporal correlated inputs.

We chose to use kind of scenario approach: it consists in grouping the 4 temporal inputs into a single input factor, defining a weather scenario.



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For sake of clarity, we consider

$$f: \left\{ \begin{array}{ccc} \mathbb{R}^d & \to & \mathbb{R} \\ \mathbf{x} = (\mathbf{x}_1, \dots, \mathbf{x}_d) & \mapsto & \mathbf{y} = f(\mathbf{x}) \end{array} \right.$$

Does the output *Y* vary more or less when fixing one of its input parameters?

$$V[Y|X_i = x_i]$$
, how to choose  $x_i$ ?

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The more this quantity is small, the more fixing  $X_i$  reduces the variance of Y: the input  $X_i$  is influent.

$$\xrightarrow[i=1,...d]{\text{1st order Sobol' indices}} 0 \le S_i = \frac{V[E(Y|X_i)]}{V[Y]} = 1 - \frac{E[V(Y|X_i)]}{V[Y]} \le 1$$



For sake of clarity, we consider  $f:[0,1]^d\to\mathbb{R}$ . Then, if  $\int_{[0,1]^d}f^2(\mathbf{x})d(\mathbf{x})<+\infty$ , f admits a unique decomposition

$$f_0 + \sum_{i=1}^d f_i(\mathbf{x}_i) + \sum_{1 \le i \le j \le d} f_{i,j}(\mathbf{x}_i, \mathbf{x}_j) + \ldots + f_{1,\ldots,d}(\mathbf{x}_1, \ldots, \mathbf{x}_d)$$

under the constraints

- ▶ f<sub>0</sub> constant,
- $\forall \ 1 \leq s \leq d, \ \forall \ 1 \leq i_1 < \ldots < i_s \leq d, \ \forall \ 1 \leq p \leq s$

$$\int_0^1 f_{i_1,\dots,i_s}(x_{i_1},\dots,x_{i_s})dx_{i_p}=0$$



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$$\int_0^1 f_{i_1,\dots,i_s}(x_{i_1},\dots,x_{i_s})dx_{i_p}=0$$

### Consequences:

$$\star f_0 = \int_{[0,1]^d} f(\mathbf{x}) d\mathbf{x},$$

$$\star \forall \mathbf{u} \subset \{1,\ldots,d\}, \int_{[0,1]^d} f_{\mathbf{u}}(\mathbf{x}_{\mathbf{u}}) d\mathbf{x} = 0,$$

$$\star \forall \mathbf{u} \neq \mathbf{v} \subset \{1, \dots, d\}, \int_{[0,1]^d} f_{\mathbf{u}}(\mathbf{x}_{\mathbf{u}}) f_{\mathbf{v}}(\mathbf{x}_{\mathbf{v}}) d\mathbf{x} = 0.$$



We deduce from the constraints

$$f_i(x_i) = \int_{[0,1]^{d-1}} f(x) \prod_{p \neq i} dx_p - f_0$$

$$i \neq j$$

$$f_{i,j}(x_i, x_j) = \int_{[0,1]^{d-2}} f(x) \prod_{p \neq i,j} dx_p - f_0 - f_i(x_i) - f_j(x_j)$$

▶ ...

We deduce from the constraints

• 
$$f_i(x_i) = \int_{[0,1]^{d-1}} f(x) \prod_{p \neq i} dx_p - f_0$$

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$$f_{i,j}(x_i, x_j) = \int_{[0,1]^{d-2}} f(x) \prod_{p \neq i,j} dx_p - f_0 - f_i(x_i) - f_j(x_j)$$

**•** ...

Or equivalently, for  $X \sim \mathcal{U}\left([0,1]^d\right)$ ,

$$Y = f(X) = f_0 + \sum_{i=1}^d f_i(X_i) + \sum_{1 \le i \le j \le d} f_{i,j}(X_i, X_j) + \dots + f_{1,\dots,d}(X_1, \dots, X_d)$$

with

$$f_i(\mathbf{x}_i) = E[Y|X_i = \mathbf{x}_i] - E[Y],$$

▶ 
$$i \neq j$$
,  $f_{i,j}(x_i, x_j) = E[Y|X_i, X_j = x_i, x_j] - E[Y|X_i = x_i] - E[Y|X_j = x_j] + E[Y]$ ,

Variance decomposition:

$$V[Y] = \sum_{i=1}^{d} V[f_i(X_i)] + ... + V[f_{1,...,d}(X_1,...,X_d)]$$

Sobol' indices:

$$\forall i = 1, \dots, d S_i = \frac{V[f_i(X_i)]}{V[Y]} = \frac{V[\mathbb{E}[Y|X_i]]}{V[Y]}$$

$$\forall i \neq j \ S_{i,j} = \frac{V[f_{i,j}(X_i,X_j)]}{V[Y]} = \frac{V[\mathbb{E}[Y|X_i,X_j]] - V[\mathbb{E}[Y|X_i]] - V[\mathbb{E}[Y|X_j]]}{V[Y]}$$

We have

$$1 = \sum_{i=1}^{d} S_i + \sum_{i \neq i} S_{i,j} + \ldots + S_{1,\ldots,d}$$

Factors Prioritization (FP): which factor should one try to determine first in order to have the largest expected reduction in the variance of the model output?  $\longrightarrow$  first order Sobol' indices do the job.



#### Total Sobol' indices:

$$i = 1, \dots, d$$
  $S_i^{\text{tot}} = \sum_{\mathbf{u} \subset \{1, \dots, d\}, \ \mathbf{u} \cap \{i\} \neq \emptyset} S_{\mathbf{u}}$ 

Factors Fixing (FF): which input factors can be fixed, anywhere in their range of variation, without sensibly affecting a specific output of interest? —> total Sobol' indices do the job.

We have:

$$S_{i}^{\mathsf{tot}} = \frac{E\left[V\left[Y|\mathbf{X}_{-i}\right]\right]}{V\left[Y\right]} = 1 - \frac{V\left[E\left[Y|\mathbf{X}_{-i}\right]\right]}{V\left[Y\right]}$$

with 
$$X_{-i} = (X_1, \dots, X_{i-1}, X_{i+1}, \dots, X_d)$$
.



### Variance based sensitivity analysis, general framework

We consider

$$f: \left\{ \begin{array}{ccc} \mathbb{R}^d & \to & \mathbb{R} \\ \mathbf{x} = (x_1, \dots, x_d) & \mapsto & y = f(\mathbf{x}) \end{array} \right.$$

P(dx) not necessarily equal to  $P_1(dx_1) \dots P_d(dx_d)$ .

Let, 
$$\forall i = 1, ..., d$$
,  $F_{X_i}(\cdot) = P(X_i \leq \cdot)$  and 
$$\forall \mathbf{x} = (\mathbf{x}_1, ..., \mathbf{x}_d) \in \mathbb{R}^d, \ F_{\mathbf{X}}(\mathbf{x}) = P(X_1 \leq \mathbf{x}_1, ..., X_d \leq \mathbf{x}_d).$$

Sklar's Theorem 
$$F_X(x) = C(F_{X_1}(x_1), \dots, F_{X_d}(x_d)).$$

If the  $F_{X_i}$  are continuous, then the copula C is unique.



### Variance based sensitivity analysis, general framework

We can still define Sobol' indices as:

$$\forall i = 1, \dots, d \ S_i = \frac{V[\mathbb{E}[Y|X_i]]}{V[Y]}$$

$$\forall i \neq j \ S_{i,j} = \frac{V[\mathbb{E}[Y|X_i,X_j]] - V[\mathbb{E}[Y|X_i]] - V[\mathbb{E}[Y|X_j]]}{V[Y]}$$
...

We do not have anymore 
$$1 = \sum_{i=1}^{d} S_i + \sum_{i \neq j} S_{i,j} + \ldots + S_{1,\ldots,d}$$

We do not have anymore

$$\sum_{\mathbf{u}\subset\{1,\dots,d\},\ \mathbf{u}\cap\{i\}\neq\emptyset} S_{\mathbf{u}} = \frac{E\left[V\left[Y|\mathbf{X}_{-i}\right]\right]}{V\left[Y\right]}$$

FP and FF are not easy tasks anymore.



# An alternative, the Shapley effects Shapley paradigm

Let team  $\mathbf{u} \subseteq \{1,2,...,d\}$  create value  $\mathrm{val}(\mathbf{u})$ . Let  $\mathbf{u} \mapsto \mathrm{val}(\mathbf{u})$ , with  $\mathrm{val}(\emptyset) = 0$ , be the characteristic function of the game. The total value of the game is  $\mathrm{val}(\{1,2,...,d\})$ . We attribute  $\phi_i$  of this to  $i \in \{1,2,...,d\}$ .

### Shapley axioms

- Efficiency  $\sum_{i=1}^{d} \phi_i = \operatorname{val}(\{1,\ldots,d\}).$
- ▶ Dummy  $[\forall \mathbf{u}, \operatorname{val}(\mathbf{u} \cup \{i\}) = \operatorname{val}(\mathbf{u})] \Rightarrow \phi_i = 0.$
- Symmetry  $[\forall \mathbf{u} \text{ such that } \mathbf{u} \cap \{i,j\} = \emptyset, \text{ val}(\mathbf{u} \cup \{i\}) = \text{val}(\mathbf{u} \cup \{j\})]$  $\Rightarrow \phi_i = \phi_i.$
- Additivity If games with characteristic functions val, val' correspond to  $\phi_i$ ,  $\phi_i'$ , then the game with characteristic function val + val' corresponds to  $\phi_i + \phi_i'$ .

[Shapley, 1953] shows there is a unique solution to these axioms.



### An alternative, the Shapley effects

The solution is:

$$\phi_i = rac{1}{d} \sum_{oldsymbol{u} \subseteq -\{i\}} inom{d-1}{|oldsymbol{u}|}^{-1} ig( ext{val}( ext{u}+ ext{i}) - ext{val}( ext{u}) ig)$$

Let variables  $x_1, x_2, \ldots, x_d$  be team members trying to explain f. The value of any subset u is how much can be explained by  $x_u$ .

The Shapley effects [Owen, 2014] are the  $\phi_i$ s corresponding to the value function  $\mathbf{u} \mapsto \sum_{\mathbf{v} \subseteq \mathbf{u}} S_{\mathbf{v}} = \frac{V[E[Y|\mathbf{X}_{\mathbf{u}}]]}{V[Y]}$ .

If we define the value function as  $\boldsymbol{u} \mapsto \frac{E[V[Y|X_u]]}{V[Y]}$  instead of  $\boldsymbol{u} \mapsto \frac{V[E[Y|X_u]]}{V[Y]}$ , we also get the Shapley effects.



# An alternative, the Shapley effects

Independent framework: 
$$\forall i = 1, ..., d$$
,  $\phi_i = \sum_{\boldsymbol{u}: i \in \boldsymbol{u}} \frac{1}{|\boldsymbol{u}|} S_{\boldsymbol{u}}$ 

We also have:  $\forall i = 1, ..., d$ ,  $0 \le S_i \le \phi_i \le S_i^{\text{tot}} \le 1$  and  $\sum_{i=1}^{d} \phi_i = 1$ .



### An alternative, the Shapley effects

Independent framework: 
$$\forall i = 1, \ldots, d$$
,  $\phi_i = \sum_{\boldsymbol{u}: i \in \boldsymbol{u}} \frac{1}{|\boldsymbol{u}|} \mathcal{S}_{\boldsymbol{u}}$ 

We also have:  $\forall i = 1, ..., d$ ,  $0 \le S_i \le \phi_i \le S_i^{\text{tot}} \le 1$  and  $\sum_{i=1}^{d} \phi_i = 1$ .

### Dependent framework:

In this framework, it is usual to define first order and total Sobol' indices as

$$S_{i} = \frac{V\left[E\left[Y|\mathbf{X}_{i}\right]\right]}{V\left[Y\right]}$$

$$S_{i}^{\text{tot}} = \frac{E\left[V\left[Y|\mathbf{X}_{-i}\right]\right]}{V\left[Y\right]}$$

We still have  $0 \le \phi_i \le 1$  and  $\sum_{i=1}^d \phi_i = 1$ We do not necessarily have  $S_i \le \phi_i \le S_i^{\text{tot}}$ 

### Some properties

▶ Property 1 Let  $Y = f(X) = f(X_1, ..., X_d)$ , with  $X_1 = h(X_2)$  and  $X_2 = h^{-1}(X_1)$  with probability one. Then  $\phi_1 = \phi_2$ .

Property 2 Let  $Y = f(\mathbf{X})$  and for i = 1, ..., d,  $X_i = \tau_i^{-1}(Z_i)$ . We define  $g(\mathbf{Z}) = f(\tau_1^{-1}(Z_1), ..., \tau_d^{-1}(Z_d))$ . Let  $\phi_i'$  be the Shapley importance of  $Z_i$  as a predictor of  $Y' = g(\mathbf{Z})$ . Then, for all i = 1, ..., d,  $\phi_i' = \phi_i$ .



[Owen and Prieur, 2017, looss and Prieur, 2017]

Gaussian framework, affine model, d = 2

We consider  $\mathbf{X} \sim \mathcal{N}_2(\boldsymbol{\mu}, \boldsymbol{\Sigma})$  and  $\mathbf{Y} = \beta_0 + \boldsymbol{\beta}^\mathsf{T} \mathbf{X}$ , with

$$\boldsymbol{\mu} = \begin{pmatrix} \mu_1 \\ \mu_2 \end{pmatrix}, \boldsymbol{\beta} = \begin{pmatrix} \beta_1 \\ \beta_2 \end{pmatrix}, \boldsymbol{\Sigma} = \begin{pmatrix} \sigma_1^2 & \rho \sigma_1 \sigma_2 \\ \rho \sigma_1 \sigma_2 & \sigma_2^2 \end{pmatrix}, \rho \in [-1, 1], \sigma_i > 0.$$

We have  $\sigma^2 = V[Y] = \beta_1^2 \sigma_1^2 + 2\rho \beta_1 \beta_2 \sigma_1 \sigma_2 + \beta_2^2 \sigma_2^2$  and

$$\sigma^2 Sh_1 = \beta_1^2 \sigma_1^2 (1 - \frac{\rho^2}{2}) + \rho \beta_1 \beta_2 \sigma_1 \sigma_2 + \beta_2^2 \sigma_2^2 \frac{\rho^2}{2},$$

$$\sigma^2 Sh_2 = \beta_2^2 \sigma_2^2 (1 - \frac{\rho^2}{2}) + \rho \beta_1 \beta_2 \sigma_1 \sigma_2 + \beta_1^2 \sigma_1^2 \frac{\rho^2}{2}.$$



Farlie-Gumbel-Morgenstern copula, uniform marginals, affine model, d=2

We consider  $Y = \beta_0 + \boldsymbol{\beta}^T \mathbf{X}$ , with

- ▶ marginal distributions  $X_i \sim \mathcal{U}([0,1]^2)$  for i = 1, 2,
- ▶ joint probability density function  $c_{\theta}(\mathbf{x_1}, \mathbf{x_2}) = 1 + \theta(1 2\mathbf{x_1})(1 2\mathbf{x_2}), -1 \le \theta \le 1.$

We then have

• 
$$\sigma_i^2 = V[X_i] = 1/12$$
,

We also have

We get once more

$$\sigma^{2}Sh_{1} = \beta_{1}^{2}\sigma_{1}^{2}(1-\frac{\rho^{2}}{2}) + \rho\beta_{1}\beta_{2}\sigma_{1}\sigma_{2} + \beta_{2}^{2}\sigma_{2}^{2}\frac{\rho^{2}}{2},$$
  
$$\sigma^{2}Sh_{2} = \beta_{2}^{2}\sigma_{2}^{2}(1-\frac{\rho^{2}}{2}) + \rho\beta_{1}\beta_{2}\sigma_{1}\sigma_{2} + \beta_{1}^{2}\sigma_{1}^{2}\frac{\rho^{2}}{2}.$$

That is

$$Sh_1 = \frac{1}{2} \left( 1 + \left( 1 - \frac{\theta^2}{9} \right) \frac{\beta_1^2 - \beta_2^2}{12\sigma^2} \right),$$

$$Sh_2 = \frac{1}{2} \left( 1 + \left( 1 - \frac{\theta^2}{9} \right) \frac{\beta_2^2 - \beta_1^2}{12\sigma^2} \right),$$

with 
$$\sigma^2 = \frac{\beta_1^2 + \beta_2^2}{12} + \beta_1 \beta_2 \frac{\theta}{18}$$
.



Gaussian framework, exponential model, d=2

Let 
$$f(\mathbf{X}) = \exp(\mathbf{X}^\mathsf{T} \boldsymbol{\beta})$$
 for  $\mathbf{X}, \boldsymbol{\beta} \in \mathbb{R}^2$  and  $\mathbf{X} \sim \mathcal{N}(\mathbf{0}, \boldsymbol{\Sigma})$ , for  $\boldsymbol{\Sigma} = \begin{pmatrix} 1 & \rho \\ \rho & 1 \end{pmatrix}$ . Then

$$\textit{Sh}_1 = \frac{1}{2} \bigg( 1 + \frac{e^{(\beta_1 + \beta_2 \rho)^2} - e^{(\beta_2 + \beta_1 \rho)^2}}{e^{\beta_1^2 + \beta_2^2 + 2\rho\beta_1\beta_2} - 1} \bigg)$$

$$Sh_2 = rac{1}{2} \left( 1 + rac{e^{(eta_2 + eta_1 
ho)^2} - e^{(eta_1 + eta_2 
ho)^2}}{e^{eta_1^2 + eta_2^2 + 2
hoeta_1eta_2} - 1} 
ight)$$

From Property 2, we can extend the result to unnormalized X1 and  $X_2$ .



#### General bivariate framework

We consider Y = f(X) with finite variance  $\sigma^2 > 0$ . We have

$$\begin{aligned} \textit{Sh}_1 &=& \frac{1}{2} \left( 1 + \frac{V[E[\textcolor{red}{Y}|\textcolor{blue}{X_1}]] - V[E[\textcolor{red}{Y}|\textcolor{blue}{X_2}]]}{\sigma^2} \right) \,, \end{aligned}$$

$$Sh_2 = \frac{1}{2} \left( 1 + \frac{V[E[Y|X_2]] - V[E[Y|X_2]]}{\sigma^2} \right) \cdot$$



Gaussian framework, model with a second order interaction, d = 3

We consider  $Y=X_1+X_2X_3$  with  $old X\sim \mathcal N_2(old 0, old \Sigma)$  and

$$\boldsymbol{\Sigma} = \begin{pmatrix} \sigma_1^2 & 0 & \rho \sigma_1 \sigma_3 \\ 0 & \sigma_2^2 & 0 \\ \rho \sigma_1 \sigma_3 & 0 & \sigma_3^2 \end{pmatrix}, \rho \in [-1, 1], \sigma_i > 0.$$

We have  $\sigma^2 = V[Y] = \sigma_1^2 + \sigma_2^2 \sigma_3^2$  and



General model with a block-additive structure, d = 3

We consider  $Y = g(X_1, X_2) + h(X_3)$ . We assume that all the three inputs have a finite variance and that  $X_3$  is independent from  $(X_2, X_3)$ .

We then have

• 
$$S_3 = Sh_3 = S_3^{\text{tot}}$$

• for 
$$i = 1, 2,$$

$$\begin{aligned} \left[ S_i \le Sh_i \right] &\Leftrightarrow \left[ Sh_i \le S_i^{\text{tot}} \right] \\ &\Leftrightarrow \left[ \frac{V[E[Y|X_1]] + V[E[Y|X_2]]}{2} \le \frac{V[E[Y|X_1, X_2]]}{2} \right] \end{aligned}$$



### What about algorithms?

Algorithms to compute Shapley effects [Castro et al., 2009] are based on the value function  $\boldsymbol{u} \mapsto \frac{E[V[Y|\mathbf{X}_u]]}{V[Y]}$ . Note that

$$Sh_i = rac{1}{d\,!} \sum_{\pi \in \Pi(\{1,\ldots,d\})} \left( \operatorname{val}(P_i(\pi) \cup \{i\})) - \operatorname{val}(P_i(\pi)) \right)$$

with  $\Pi(\{1,\ldots,d\})$  the set of all possible permutations of the inputs and for a permutation  $\pi \in \Pi(\{1,\ldots,d\})$ , the set  $P_i(\pi)$  is defined as the inputs that precede input i in  $\pi$ .

Exact permutation algo. (moderate d) all possible permutations are covered.

Random permutation algo. (d >> 1) it randomly sample permutations of the inputs.



### What about algorithms?

In [Song et al., 2016],  $val(\textbf{\textit{u}}) \rightarrow \widehat{val}(\textbf{\textit{u}})$ .

For each iteration of the loop on the inputs' permutations, a conditional variance expectation must be computed.

The cost C of these algorithms is the following:

$$C = N_v + m(d-1)N_0N_i$$

with  $N_V$  the sample size for the variance computation,  $N_0$  the outer loop size for the expectation,  $N_i$  the inner loop size for the conditional variance and m the number of permutations according to the selected method.

Bootstrap confidence intervals can be computed. A costly model can be replaced by a metamodel. [looss and Prieur, 2017, Benoumechiara and Elie-Dit-Cosaque, 2018]



### An industrial application

Industrial problem: ultrasonic non-destructive control of a weld containing manufacturing defect

The heterogeneous and anisotropic weld's structure is represented by a simplified model consisting of a partition of 7 equivalent homogeneous regions with a specific grain orientation.





Metallographic picture (left)
Description of the weld in 7 homogeneous domains (right)



Inspection configuration

Input parameters: 11 scalar inputs (4 elastic coefficients and 7 orientations).

Scalar output: the amplitude of the defect echoes resulting from an ultrasonic inspection.

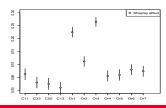
### An industrial application

Probabilistic model: all the inputs are modeled by Gaussian distributions, there are dependences between the orientations.

$$(Or_1,\ldots,Or_7)^{\mathsf{T}}\sim \mathcal{N}_7(\mu,\Sigma)$$
 [Rupin et al., 2014, Moysan et al., 2003]

Shapley effects for the ultrasonic non-destructive control application. The vertical bars represent the 95%-confidence intervals of each effect.

Algorithm's parameters:  $m = 10^4$ ,  $N_i = 3$ ,  $N_0 = 1$ ,  $N_v = 10^4$ , total cost  $C = 3 \times 10^5$  metamodel evaluations.



### Conclusion, perspectives

Conclusion: Shapley effects present an alternative to allocate parts of variance in the dependent framework. There exist algorithms to estimate these indices.

### Perspectives

- Can we propose goal-oriented Shapley effects?
- What are the theoretical finite sample properties of both algorithms?
- How can Shapley effects be related to gradient-based measures of sensitivity?
- **•** . . .



#### **Thanks**

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