Model Uncertainty and Uncertainty Quantification

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Rumsfeld's "Known Unknowns" versus "Unknown Unknowns" 🛓 🦻 🧟

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Each model corresponds to a parametric (although possibly infinite-dimensional) distribution of the data Y:

$$p_m(\mathbf{y} \mid \boldsymbol{\theta}_m, \mathbf{x}) = p(\mathbf{y} \mid \boldsymbol{\theta}_m, \mathcal{M}_m, \mathbf{x})$$

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 Objective: Obtain predictive distributions or summaries at inputs x\*

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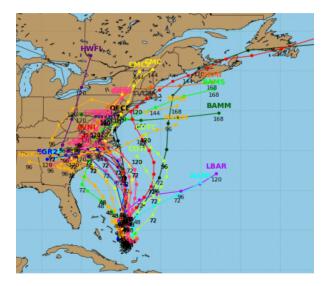
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WLOG drop dependence on inputs,  $p(\mathbf{y}^* \mid \mathbf{y})$ 

# Multiple Models



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Bayesian Perspectives on Model Uncertainty

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 $\mathcal{M}$ -Open we know the true model  $\mathcal{M}_T$  is not in M, but we cannot specify the explicit form  $p(y^* | \mathbf{y})$  because it is too difficult conceptually or computationally, we lack time to do so, or do not have the expertise, etc.

Bernardo & Smith (1994), Clyde & Iversen (2013)

# ${\sf Predictive \ Distributions \ under \ } {\cal M}{\text{-}}{\sf Closed}$

► A Bayesian would assign a prior probability, p(M<sub>m</sub>), representing their belief that each model M<sub>m</sub> is the true model.

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▶ Bayes Theorem: posterior probability of each model p(M<sub>m</sub> | Y)

$$p(\mathcal{M}_m \mid \mathbf{Y}) = \frac{p(\mathbf{Y} \mid \mathcal{M}_m) p(\mathcal{M}_m)}{\sum_{m \in M} p(\mathbf{y} \mid \mathcal{M}_m) p(\mathcal{M}_m)}, \quad m \in M$$
  
where  $p(\mathbf{Y} \mid \mathcal{M}_m) = \int p(\mathbf{Y} \mid \boldsymbol{\theta}_m, \mathcal{M}_m) p(\boldsymbol{\theta}_m \mid \mathcal{M}_m) d\boldsymbol{\theta}_m$ 

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Predictive distribution

$$p(\mathbf{y}^*|\mathbf{y}) = \sum_{m \in \mathcal{M}} p(\mathbf{y}^*|\mathcal{M}_m, \mathbf{y}) p(\mathcal{M}_m|\mathbf{y})$$
$$= \sum_{m \in \mathcal{M}} \left[ \int p(\mathbf{y}^*|\mathcal{M}_m, \theta_m, \mathbf{y}) p(\theta_m|\mathbf{y}, \mathcal{M}_m) \, d\theta_m \right] p(\mathcal{M}_m|\mathbf{y})$$

#### Estimation and Prediction

Consider the decision problem of estimation/prediction under squared error loss

$$u(Y^*,a) = -(Y^*-a)^2$$

where a is a possible action (u is utility or negative loss) and  $Y^*$  is an unknown.

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From a Bayesian perspective, the solution is to find the action that maximizes expected utility given the observed data  $\mathbf{Y}$ :

$$\mathsf{E}_{\mathbf{Y}^*|\mathbf{Y}}[u(\mathbf{Y}^*,a)] = -\int (y^*-a)^2 p(\mathbf{y}^*|\mathbf{y}) d\mathbf{y}^*$$

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where the expectation is taken with respect to the predictive distribution of  $\mathbf{Y}^*$  given the observed data  $\mathbf{y}$ .

 Under the *M*-closed perspective, optimal solution for prediction is Bayesian Model Averaging

$$a^* = \mathsf{E}_{\mathbf{Y}^*}[\mathbf{Y}^* \mid \mathbf{Y}] = \sum_{m \in \mathcal{M}} p(\mathcal{M}_m \mid \mathbf{Y}) \hat{Y}^*_{\mathcal{M}_m}$$

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- ► Use joint posterior distribution on θ | M and M to obtain prediction intervals
- Full propagation of all "known" uncertainties
- Extensive literature for regression and generalized linear models [Hoeting et al 1999, Clyde & George 2004, Bayarri et al 2012] with invariant priors/Spike & Slab + software
- more complex models via RJ-MCMC, SMC, ABC

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True Model 
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BMA model weights converge to 1 for the model that is "closest" to true model in Kullback-Leibler divergence

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Other model ensembles ?

#### Combining Models as a Decision Problem

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▶ Focus on *M*-Open case

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Guterriez-Pena & Walker approximation to a (limiting)
Dirichlet process model for estimating unknown distribution F for M<sub>T</sub>

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Equivalent representation (Lagrangian):

$$-\frac{1}{J}\sum_{j=1}^{J}\left(Y_{j}-\sum_{m\in\mathcal{M}}w_{m}\hat{Y}_{(-j),\mathcal{M}_{m}}\right)^{2}-\lambda_{0}\left(\sum_{m}^{M}w_{m}-1\right)+\sum_{m}^{M}\lambda_{m}w_{m}$$

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Super-Learners! h2oEnsemble

Predict short *vs.* long-term survival given primary tumor's molecular phenotype.

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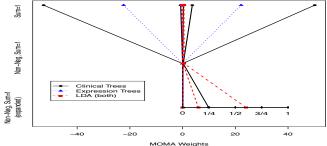
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### **MOMA** Weights



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### Validation Experiment

 5-fold cross validation; 5 splits of data into two groups: Training Y and Validation Y\*

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- Compute classification accuracy over 5 Splits

## MOMA with Sum-to-1 Constraint

	set1	set2	set3	set4	set5
clin1	53.08	-4.43	-0.01	-24.41	15.94
clin2	-79.92	-5.16	0.90	0.80	-4.63
clin3	-1.25	-0.24	-0.90	-0.01	5.35
clin4	27.36	10.14	-0.33	23.73	-17.24
clin5	1.13	0.27	0.27	0.36	0.55
tree1	-0.05	-0.55	-2.92	0.03	27.93
tree2	-0.12	-0.07	-3.21	-0.62	0.63
tree3	0.51	0.53	0.15	0.48	-3.35
tree4	-0.28	0.22	6.26	-0.04	-24.10
lda100.P1	-0.40	0.04	-0.01	0.02	-0.11
lda100.P2	0.44	-0.02	0.53	-0.06	-0.07
lda200.P1	0.30	0.17	-0.32	0.09	-0.03
Ida200.P2	0.21	0.08	0.60	0.63	0.12
Accuracy	0.64	0.64	0.46	0.73	0.60

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## MOMA with Non-negativity Constraint

	set1	set2	set3	set4	set5
clin1	0.00	0.07	0.00	0.00	0.00
clin2	0.00	0.00	0.00	0.00	0.00
clin3	0.00	0.00	0.00	0.00	0.00
clin4	0.00	0.11	0.00	0.00	0.00
clin5	0.30	0.17	0.07	0.41	0.00
tree1	0.00	0.00	0.00	0.00	0.77
tree2	0.00	0.00	0.00	0.00	0.21
tree3	0.23	0.44	0.21	0.00	0.01
tree4	0.00	0.00	0.00	0.00	0.01
lda100.P1	0.00	0.00	0.00	0.00	0.00
lda100.P2	0.22	0.00	0.30	0.00	0.00
lda200.P1	0.00	0.00	0.00	0.00	0.00
lda200.P2	0.26	0.21	0.41	0.58	0.00
Accuracy	0.82	0.73	0.55	0.73	0.60

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$$\arg\max_{\mathbf{w},\sigma^2}\sum_i \log(\sum_m^M w_m p(y_i^* \mid \mathbf{y}, \mathcal{M}_m, \sigma^2))$$

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Mixture Models and Mixtures of Experts

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